

# Do Genetic Susceptibility Variants Associate with Disease Severity in Early Active Rheumatoid Arthritis?

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**ABSTRACT. Objective.** Genetic variants affect both the development and severity of rheumatoid arthritis (RA). Recent studies have expanded the number of RA susceptibility variants. We tested the hypothesis that these associated with disease severity in a clinical trial cohort of patients with early, active RA.

**Methods.** We evaluated 524 patients with RA enrolled in the Combination Anti-Rheumatic Drugs in Early RA (CARDERA) trials. We tested validated susceptibility variants — 69 single-nucleotide polymorphisms (SNP), 15 HLA-DRB1 alleles, and amino acid polymorphisms in 6 HLA molecule positions — for their associations with progression in Larsen scoring, 28-joint Disease Activity Scores, and Health Assessment Questionnaire (HAQ) scores over 2 years using linear mixed-effects and latent growth curve models.

**Results.** HLA variants were associated with joint destruction. The \*04:01 SNP (rs660895,  $p = 0.0003$ ), \*04:01 allele ( $p = 0.0002$ ), and HLA-DR $\beta$ 1 amino acids histidine at position 13 ( $p = 0.0005$ ) and valine at position 11 ( $p = 0.0012$ ) significantly associated with radiological progression. This association was only significant in anticitrullinated protein antibody (ACPA)-positive patients, suggesting that while their effects were not mediated by ACPA, they only predicted joint damage in ACPA-positive RA. Non-HLA variants did not associate with radiograph damage (assessed individually and cumulatively as a weighted genetic risk score). Two SNP — rs11889341 (*STAT4*,  $p = 0.0001$ ) and rs653178 (*SH2B3-PTPN11*,  $p = 0.0004$ ) — associated with HAQ scores over 6–24 months.

**Conclusion.** HLA susceptibility variants play an important role in determining radiological progression in early, active ACPA-positive RA. Genome-wide and HLA-wide analyses across large populations are required to better characterize the genetic architecture of radiological progression in RA. (First Release May 15 2015; J Rheumatol 2015;42:1131–40; doi:10.3899/jrheum.141211)

## Key Indexing Terms:

RHEUMATOID ARTHRITIS  
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OUTCOME MEASURES  
HLA ANTIGENS

Identifying biomarkers that predict rheumatoid arthritis (RA) prognosis at first presentation is an important research goal. They could guide decisions on initial treatment intensity, facilitating stratified medicine. The progression of radiological damage that characterizes severe RA is 45% to 58%

heritable<sup>1</sup>. This suggests that genetic variants are potential prognostic markers. Current knowledge of the genetic architecture of radiograph progression is limited, with up to half of its heritability unexplained<sup>2</sup>.

The concept that RA susceptibility variants also influence

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disease severity is an intuitive one. Several studies have tested the hypothesis that RA risk variants associate with radiographic damage. They have several limitations. First, many associations are not replicated<sup>3,4</sup>. Second, not all risk loci have been studied (the largest analysis examined 37 independent variants)<sup>5</sup>. Third, they represent observational studies assessing patients with established RA of variable activity<sup>6,7</sup>. Because randomized controlled trials (RCT) of early, active RA<sup>8,9,10</sup> provide the evidence base for intensive therapy, for genetic variants to guide treatment intensity decisions, they also require evaluation in patients with early, active RA.

Metaanalyses of genome-wide association studies (GWAS) have expanded the number of RA susceptibility variants. These span single-nucleotide polymorphisms (SNP), HLA-DRB1 alleles, and HLA molecule amino acid polymorphisms<sup>11,12</sup>. We tested the hypothesis that they were also associated with disease severity by evaluating their relationship with radiographic progression, disease activity, and disability over 2 years in an RCT cohort of patients with early, active RA.

## MATERIALS AND METHODS

**Subjects.** We studied the combined data of patients in two 24-month trials of Combination Anti-Rheumatic Drugs in Early RA (CARDERA-1 and CARDERA-2) that recruited 467 and 159 patients with RA, respectively, from 42 rheumatology units in England<sup>13,14</sup>. Patients had less than 2 years of RA duration and active disease, defined as 3 of (1)  $\geq 3$  swollen joints, (2)  $\geq 6$  tender joints, (3)  $\geq 45$ -min morning stiffness, or (4) erythrocyte sedimentation rate (ESR)  $\geq 28$  mm/h.

CARDERA-1 patients were randomized to 4 treatment groups: (1) methotrexate (MTX), (2) MTX and cyclosporine, (3) MTX and prednisolone, and (4) MTX, cyclosporine, and prednisolone. CARDERA-2 patients were randomized to 2 treatment groups: MTX, and MTX and anakinra.

**Disease outcomes.** In CARDERA-1, modified Larsen scores, Disease Activity Scores on a 28-joint count (DAS28), and Health Assessment Questionnaire (HAQ) scores were recorded 6-monthly<sup>13</sup>. In CARDERA-2, modified Larsen scores were recorded 12-monthly, and DAS28 and HAQ scores were measured at 0, 6, 12, and 24 months. In both studies, the DAS28-ESR was used. Mean scores for each outcome at each timepoint are shown in Figure 1.

**Genotyping.** Archived DNA was genotyped on the Illumina ImmunoChip<sup>15</sup>. Quality control (QC) procedures are described in Supplementary Table 1 (available online at [jrheum.org](http://jrheum.org)). From 560 genotyped individuals, 524 were included in our analysis. From 196,524 genetic markers, 138,873 passed QC procedures (genotyping rate 0.999).

Classical HLA-DRB1 alleles and amino acid polymorphisms in HLA-DR $\beta$ 1, HLA-B, and HLA-DP $\beta$ 1 were imputed using SNP2HLA (Beagle imputation method<sup>16</sup>, reference data from Type 1 Diabetes Genetics Consortium panel<sup>17</sup>). Imputed allele/amino acid dosages were analyzed. These probabilistic representations of the number of predicted alleles correlate better with true genotypes when compared with “best-guess” genotypes<sup>16</sup>. Imputed markers with INFO scores  $< 0.5$  were excluded<sup>18</sup>.

**SNP.** We evaluated validated European ancestry RA susceptibility SNP from the recent GWAS metaanalysis of European and Asian populations by Okada, *et al*<sup>11</sup>. To maximize study power, we evaluated only loci attaining genome-wide significance in Europeans. This omitted small effect size loci (requiring the larger transethnic sample size to achieve significance) and loci that had a stronger effect in Asians.

Of 77 potential susceptibility SNP, 69 were available in CARDERA (including 31 proxy SNP,  $r^2 > 0.8$ ; Supplementary Table 2, available online at [jrheum.org](http://jrheum.org)). As neither the validated HLA region SNP (rs9268839) in the Okada, *et al* metaanalysis nor any proxy SNP were present in CARDERA, we used the validated HLA region SNP from the previous metaanalysis by Eyre, *et al* (rs660895)<sup>19</sup>. This ensured that a validated HLA susceptibility SNP was tested. Missing data in 9 SNP (7 SNP missing in 1 patient, 1 SNP in 4 patients, 1 SNP in 21 patients) were imputed by assigning them the expected allele count (twice the allele frequency)<sup>20,21</sup>.

**HLA alleles and amino acids.** Four-digit HLA-DRB1 alleles attaining  $P_{\text{GWAS}} < 5 \times 10^{-8}$  in Europeans from the RA metaanalysis of association in the HLA region<sup>12</sup> were evaluated. Of 17 potential alleles, 2 were unavailable in CARDERA (\*11:03, \*14:04) because of low INFO scores. Amino acids at positions 11, 13, 71, and 74 in HLA-DR $\beta$ 1; position 9 in HLA-B; and position 9 in HLA-DP $\beta$ 1 were assessed. In the HLA metaanalysis, these explained most of the region’s association with RA<sup>12</sup>. Positions 11 and 13 in HLA-DR $\beta$ 1 were both included because their strong linkage disequilibrium (LD) meant causality could not be assigned<sup>12</sup>.

**Associations with radiological progression.** The association between Larsen score progression and each genetic variant was examined using a linear mixed-effects (LME) model. This assessed all repeated Larsen scores simultaneously, and through incorporating correlated random effects, accounted for within-individual correlations in Larsen scores over time. This repeated measures modelling approach was preferable to evaluating associations with the 24-month change in Larsen scores ( $\Delta$ Larsen) because it (1) better evaluated within-individual changes over time<sup>22</sup>, (2) optimized the analysis power<sup>22</sup>, and (3) could accommodate missing radiographic data by assuming these observations were missing-at-random<sup>23</sup>.

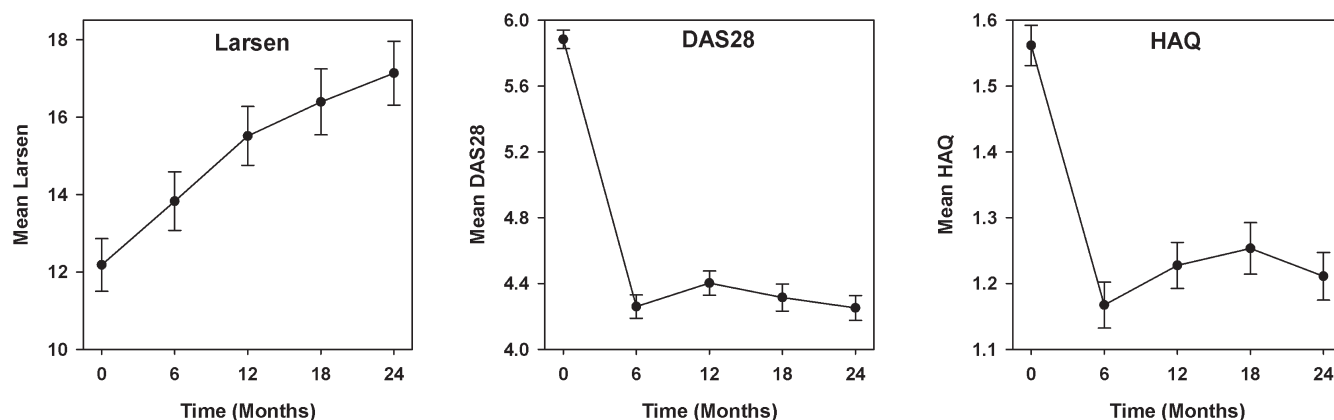


Figure 1. Mean Larsen, DAS28, and HAQ scores over 2 years. Standard error bars for each score shown. DAS28: Disease Activity Score at 28 joints; HAQ: Health Assessment Questionnaire.

Larsen scores were log-transformed with a constant added (because of zero unit scores) to approximate a normal distribution.  $\text{Log}_e$  (Larsen score + 1) was the response variable. Time (yrs), clinical variables, genotype, and a genotype\*time interaction term were fixed-effects predictor variables. Clinical variables (chosen using stepwise backward elimination) included treatment (coded categorically with MTX as the reference group), age, disease duration, and rheumatoid factor (RF) status. Their associations with Larsen scores are provided in Supplementary Table 3 (available online at [jrheum.org](http://jrheum.org)). Genotype was coded additively;  $\beta$  values were back-transformed to the original Larsen scale. The genotype\*time term  $\beta$  provided information on the annual increase in Larsen score per risk allele copy relative to the reference genotype;  $p$  values for this term provided information on the variant's role in radiological progression. The X-chromosome SNP (rs2075596) was assessed separately in men and women.

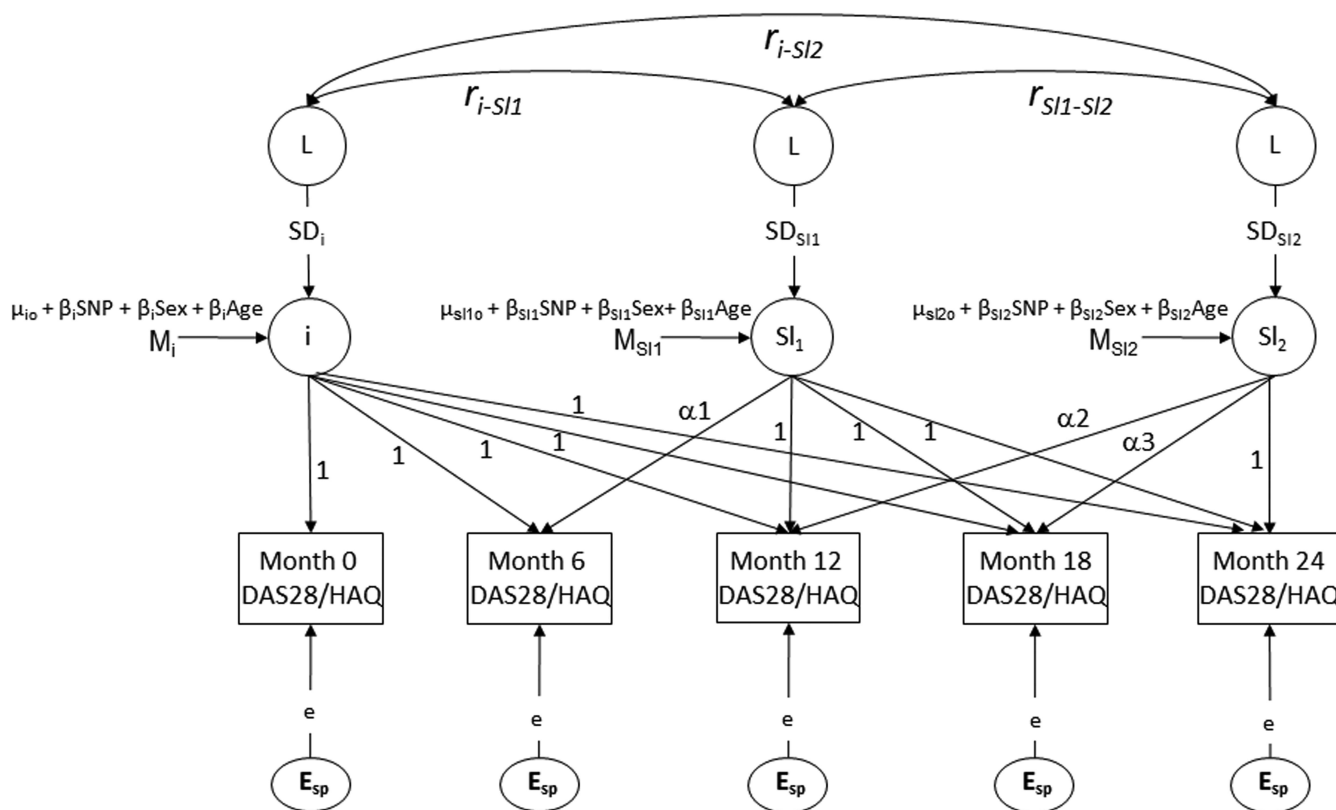
RF was used as a covariate instead of anticitrullinated protein antibodies (ACPA) because of missing ACPA data (25 patients). Significant variants were (1) reanalyzed in a model with ACPA as a covariate and (2) evaluated in ACPA-positive and ACPA-negative patients separately (omitting serology as a covariate); this established whether their effects were independent of ACPA. Examination of residuals from a model containing time and clinical variables confirmed a good model fit.

To ensure the LME model results were not biased by missing radiographic data, a supplementary approach was undertaken. This involved a 3-way ANOVA model including the  $\Delta$ Larsen as the response variable and

treatment, RF status, and genotype (coded additively) as predictor variables (chosen through stepwise backward elimination). Because the  $\Delta$ Larsen remained non-normally distributed despite log transformation, a rank-based inverse normal transformation was undertaken<sup>24</sup>. Residual plots confirmed a good model fit. The 19 individuals without 24-month radiograph data were excluded from this analysis.

**Associations with DAS28 and HAQ scores.** The associations between each variant and DAS28 and HAQ scores were assessed using latent growth curve models (GCM). In contrast to Larsen scores, longitudinal changes in DAS28/HAQ scores were nonlinear: in months 0–6, an improvement was seen; subsequently a worsening then improvement was observed (Figure 1). A 2-piece GCM (Figure 2) was therefore used, composing of an intercept, slope 1 (0 to 6–mo change in DAS28/HAQ), and slope 2 (6 to 24–mo change in DAS28/HAQ). The GCM estimated the effect of each genotype on the means of the intercept, slope 1, and slope 2, corrected for the covariates sex and age (chosen through stepwise backward elimination).

**Weighted genetic risk score (wGRS) associations with disease outcomes.** The associations between a wGRS combining susceptibility loci, and each outcome were tested. Three wGRS were evaluated: (1) a full wGRS (all SNP except the X-chromosome marker), (2) an HLA wGRS (all HLA-DRB1 alleles), and (3) a non-HLA wGRS (as per the full wGRS, but excluding the HLA SNP). The wGRS replaced the genotype term in the LME/ANOVA/GCM models.



**Figure 2.** Multiphase latent growth curve factor model for DAS28 and HAQ scores.  $i$ : intercept.  $Sl_1$ : slope 1 (identifies change in DAS28/HAQ from month 0–6).  $Sl_2$ : slope 2 (identifies change in DAS28/HAQ from month 6–24).  $E_{sp}$ : residual error terms (constrained to be equal).  $M_i$ ,  $M_{Sl1}$ , and  $M_{Sl2}$ : intercept, slope 1, and slope 2 means, respectively.  $\mu_{i0}$ ,  $\mu_{Sl10}$ , and  $\mu_{Sl20}$ : overall mean estimate of the intercept, slope 1, and slope 2 factors, respectively.  $\beta_i \text{SNP}$ ,  $\beta_{Sl1} \text{SNP}$ , and  $\beta_{Sl2} \text{SNP}$ : SNP dependent intercept, slope 1, and slope 2 means, respectively.  $\beta_i \text{Sex}$ ,  $\beta_{Sl1} \text{Sex}$ , and  $\beta_{Sl2} \text{Sex}$ : sex-dependent intercept, slope 1, and slope 2 means, respectively.  $\beta_i \text{Age}$ ,  $\beta_{Sl1} \text{Age}$ , and  $\beta_{Sl2} \text{Age}$ : age-dependent intercept, slope 1, and slope 2 means, respectively.  $r_{i-Sl1}$ ,  $r_{i-Sl2}$  are the correlations between the intercept and slopes.  $r_{Sl1-Sl2}$  the correlation between slope 1 and slope 2.  $SD_i$ ,  $SD_{Sl1}$ , and  $SD_{Sl2}$  are the SD of the latent factors. The overall variance-covariance structure and the overall means of the HAQ/DAS28 scores, which are optimized, are modeled by the variance-covariance structure and the means of the latent factors weighted by the factor loadings. DAS28: Disease Activity Score at 28 joints; HAQ: Health Assessment Questionnaire; SNP: single-nucleotide polymorphisms.



wGRS were constructed using the R package, REGENT<sup>25,26</sup>, which uses data on allelic OR from reference metaanalyses<sup>11,12</sup> to generate a genotype relative risk (GRR) score across loci within a multiplicative model. The log (GRR) represented the wGRS. SNP allelic OR were obtained from the reference metaanalysis' combined stage<sup>11</sup>.

**Significance thresholds.** For SNP, HLA allele, and amino acid association tests, Bonferroni corrected p value thresholds of 0.0007 (69 markers), 0.0033 (15 markers), and 0.0019 (27 markers), respectively, were considered significant. For the 4 HLA markers tested separately in ACPA subsets, a p value threshold of 0.0125 was used.

**Power calculations.** CARDERA had 81% power to detect an SNP accounting for 3.4% of the variance in each outcome<sup>27</sup>.

**Statistical programs.** Analyses were performed in R version 3.0.2 (LME/ANOVA), OpenMx version 1.4 (latent GCM), and PLINK version 1.07<sup>28</sup>.

**Ethical approval.** CARDERA-1 [Research Ethics Committee (REC) reference: MREC (1) 99/04] and CARDERA-2 (REC reference: MREC 02/1/089) were ethically approved. Approval was obtained to genotype archived DNA (REC reference: 11/EE/0544). All patients provided consent.

RESULTS

**Subjects.** The 524 patients (Table 1) were mostly women (69%) and ACPA-positive (70%). Their mean age was 54.7 years. Median baseline disease duration was 1 month; 40 patients (8%) had disease durations > 12 months. Their high initial mean DAS28 (5.88) confirmed active RA. Some had radiological damage (median Larsen score 6.50) at baseline, which was expected in patients with active RA with disease durations up to 2 years.

**SNP.** Only rs660895 (p = 0.0003) had a p value for an association with radiographic progression that passed the pre-defined significance threshold when tested using the LME model (Table 2). This SNP tags \*04:01; its β value of 1.07 (95% CI 1.03–1.12) indicated a 1.07-fold greater annual

increase in Larsen scores per copy of the risk (G) allele carried, relative to the reference (A) allele (Figure 3). This SNP remained significant in a model including ACPA as a covariate in place of RF (p = 0.0003).

Restricting analyses to ACPA-positive patients (Table 2) provided similar findings for this SNP (β = 1.07, p = 0.0037). No significant association was seen in ACPA-negative patients (β = 1.03, p = 0.4571), although the sample size was substantially smaller. This suggested that rs660895's effect on radiological progression was not mediated by ACPA; if rs660895 drove ACPA formation, which itself caused joint destruction, then no association between this SNP and radiographic progression would be seen in ACPA-positive patients. Results for all SNP are given in Supplementary Table 4 (available online at jrheum.org).

Using the ANOVA model to test the association between each SNP and the ΔLarsen, the 2 most strongly associated SNP — rs660895 (p = 0.0002) and rs7579944 (p = 0.0140) — were the same as those identified by the LME model. Only rs660895 was significant.

**HLA-DRB1 alleles.** Only \*04:01 (p = 0.0002) had a p value that passed the predefined significance threshold using the LME model (Table 2). Its β value of 1.10 (95% CI 1.05–1.15) indicated a 1.10-fold greater annual increase in Larsen scores per \*04:01 allele copy carried, relative to a non-carrier. This allele remained significant in a model including ACPA as a covariate instead of RF (p = 0.0006). Restricting analyses to ACPA-positive patients revealed similar findings (β = 1.08, p = 0.0105). No significant effect was seen in ACPA-negative patients (β = 1.04, p = 0.4344). Results for all alleles are given in Supplementary Table 5 (available online at jrheum.org). Using the ANOVA model to test the association between HLA-DRB1 susceptibility alleles and the ΔLarsen, \*04:01 remained significant (p = 0.0009).

**HLA amino acids.** Two amino acid polymorphisms — histidine at position 13 (His13; β = 1.07, p = 0.0005) and valine at position 11 (Val11; β = 1.07, p = 0.0012) in HLA-DRβ1 — had p values that passed the predefined significance threshold using the LME model (Table 2). Repeating the analysis with ACPA as a covariate instead of RF revealed similar results (His13 p = 0.0006, Val11 p = 0.0012). Restricting analyses to ACPA-positive patients showed similar findings; no significant association was seen in ACPA-negative patients (Table 2). Using the ANOVA model, His13 (p = 0.0002) and Val11 (p = 0.0002) in HLA-DRβ1 remained the most significant markers. Results for all amino acids are given in Supplementary Table 6 (available online at jrheum.org).

Because His13 and Val11 were in tight LD (r<sup>2</sup> = 0.943), a conditional analysis was undertaken. Conditioning on His13 eliminated the effect of Val11 (LME p = 0.5547, ANOVA p = 0.4851). Conditioning on Val11 eliminated the effect of His13 (LME p = 0.1699, ANOVA p = 0.8697). This indicated the significance at both these markers was driven by their

Table 1. CARDERA genetics cohort patient baseline data. Values are n (%) unless otherwise specified.

Characteristics		Summary Statistic
Demographic	Female	359 (69)
	Age, yrs, mean (95% CI)	54.7 (53.6–55.8)
RA-specific	RA duration, mos, median (IQR)	1.0 (0.0–4.0)
	ACPA-positive*	350 (70)
	RF-positive	352 (67)
	Larsen score, median (IQR)	6.50 (2.00–16.50)
	DAS28, mean (95% CI)	5.88 (5.77–5.99)
Treatment	HAQ, mean (95% CI)	1.56 (1.50–1.62)
	Receiving MTX	161 (31)
	Receiving MTX and cyclosporine	108 (21)
	Receiving MTX and prednisolone	102 (19)
	Receiving triple therapy	107 (20)
	Receiving MTX and anakinra	46 (9)

\* ACPA status missing in 25 patients. RA: rheumatoid arthritis; IQR: interquartile range; ACPA: anticitrullinated protein antibodies; RF: rheumatoid factor; DAS28: Disease Activity Score at 28 joints; HAQ: Health Assessment Questionnaire; MTX: methotrexate; triple therapy: MTX, cyclosporine, and prednisolone; CARDERA: Combination Anti-Rheumatic Drugs in Early RA trials.

Table 2. Genetic variants attaining a significant association with radiological progression.

SNP (Gene) or HLA-DRB1 Allele or HLA Amino Acid	All RA, n = 524		ACPA+, n = 350		ACPA-, n = 149	
	$\beta$ (95% CI)	p	$\beta$ (95% CI)	p	$\beta$ (95% CI)	p
rs660895 ( <i>HLA-DRB1</i> *04:01 tag)	1.07 (1.03–1.12)	0.0003	1.07 (1.02–1.13)	0.0037	1.03 (0.96–1.11)	0.4571
<i>HLA-DRB1</i> *04:01	1.10 (1.05–1.15)	0.0002	1.08 (1.02–1.14)	0.0105	1.04 (0.94–1.16)	0.4344
Histidine at position 13 in HLA-DR $\beta$ 1	1.07 (1.03–1.12)	0.0005	1.07 (1.02–1.13)	0.0060	1.02 (0.95–1.10)	0.5590
Valine at position 11 in HLA-DR $\beta$ 1	1.07 (1.03–1.11)	0.0012	1.07 (1.01–1.12)	0.0111	1.02 (0.95–1.10)	0.6130

p values from the linear-mixed effects model genotype\*time interaction term. ACPA status missing in 25 patients. Bonferroni-corrected p value thresholds for SNP, HLA-DRB1 alleles, and HLA amino acid polymorphisms in all RA are 0.0007, 0.0033, and 0.0019, respectively. Bonferroni-corrected p value threshold for the 4 markers tested in ACPA-positive and ACPA-negative RA is 0.0125. SNP: single-nucleotide polymorphism; RA: rheumatoid arthritis; ACPA: anti-citrullinated protein antibodies.

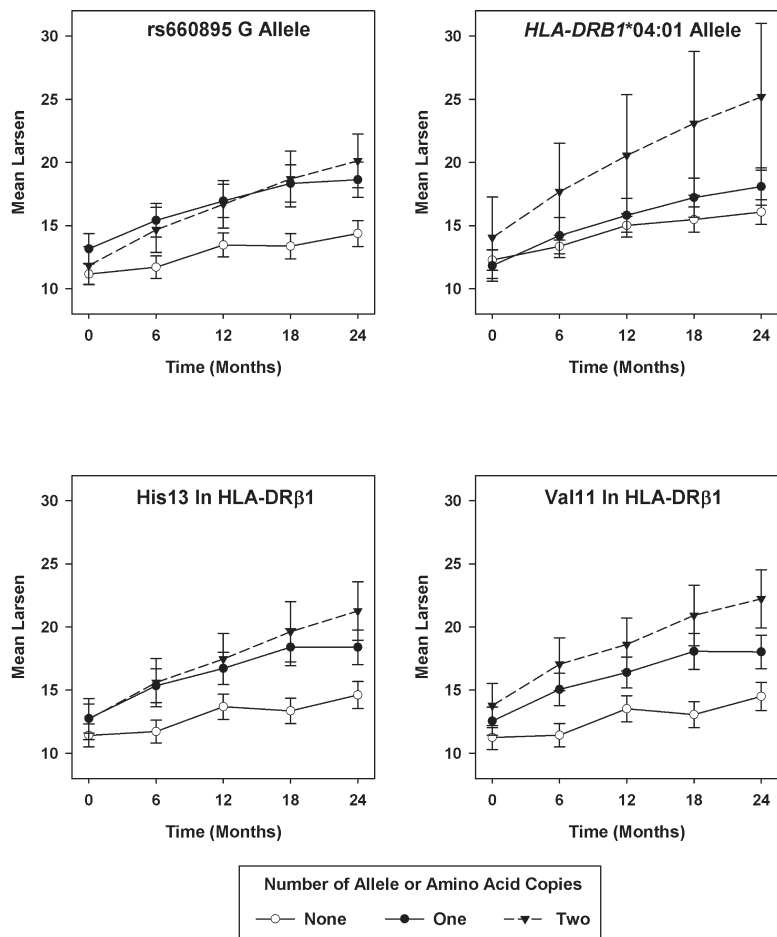


Figure 3. Significant genetic associations with Larsen score progression. Mean Larsen scores with standard error bars shown at each timepoint stratified by genotype ("best-guess" genotypes used for imputed data in this figure, but not in the statistical analysis). His13: histidine at position 13; Val11: valine at position 11.

high LD, although it was not possible to establish which variant drove their association with Larsen scores.

None of the tested polymorphisms in the shared epitope (SE) region (positions 71 and 74) were significant (Supplementary Table 6 available online at [jrheum.org](http://jrheum.org)); the lowest p value was for lysine at position 71 ( $p = 0.0293$ ). However, the SE itself (sequences QRRAA, RRRAA, and QKRAA in positions 70–74<sup>29</sup>) significantly associated with radiological progression when the classical SE alleles (\*01:01, \*01:02, \*04:01, \*04:04, \*04:05, \*04:08, \*10:01) were grouped together (LME  $p = 0.0003$ , ANOVA  $p = 0.0001$ ).

Although His13 and Val11 are not in LD with positions 71 and 74, there is a marked correlation between their presence and carrying an SE sequence. His13 and Val11 are encoded by the SE alleles \*04:01, \*04:04, \*04:05, \*04:08, and \*10:01<sup>12</sup>. A conditional analysis incorporating the SE as a modeling covariate was therefore undertaken. Conditioning on the SE eliminated the effect of His13 (LME  $p = 0.1996$ , ANOVA  $p = 0.1346$ ) and Val11 (LME  $p = 0.3593$ , ANOVA  $p = 0.1224$ ). Conditioning on His13 (LME  $p = 0.08301$ , ANOVA  $p = 0.0842$ ) and Val11 (LME  $p = 0.05345$ , ANOVA  $p = 0.1023$ ) eliminated the effect of the SE. This suggests that the association of His13, Val11, and the SE with joint destruction was driven by a high correlation between them, although it was not possible to establish which variant drove the association.

**wGRS.** The full wGRS (LME  $p = 0.0428$ ) and HLA wGRS (LME  $p < 0.0001$ ), but not the non-HLA wGRS (LME  $p = 0.6720$ ), associated with radiological progression (Supplementary Table 7, available online at [jrheum.org](http://jrheum.org)).

**Associations with disease activity.** Ten SNP, 1 HLA-DRB1 allele, and 4 HLA amino acid polymorphisms had p values  $< 0.05$  for an association with the intercept and/or slope 1 and/or slope 2 DAS28 latent growth factor means (Supplementary Table 8, available online at [jrheum.org](http://jrheum.org)). None were significant after Bonferroni correction.

The HLA wGRS associated with the intercept DAS28 latent growth factor means ( $\beta = 0.09$ ,  $p = 0.041$ ), but not those for slope 1 or slope 2 (Supplementary Table 7, available online at [jrheum.org](http://jrheum.org)). No associations were seen between the full wGRS or non-HLA wGRS and DAS28.

**Associations with disability.** From all tested variants, 2 SNP had p values for an association with HAQ score slope 2 latent growth factor means that passed the predefined significance threshold (Supplementary Table 9, available online at [jrheum.org](http://jrheum.org)). These were composed of rs11889341 in *STAT4* ( $\beta = 0.142$ ,  $p = 0.0001$ ) and rs653178 in *SH2B3-PTPN11* ( $\beta = -0.109$ ,  $p = 0.0004$ ). Although the rs11889341 RA risk (T) allele associated with greater HAQ score increases over months 6–24, mean HAQ scores were lower at all timepoints in T allele carriers (Figure 4). Similarly, although the rs653178 RA risk (C) allele associated with an improvement in HAQ scores over months 6–24, at all timepoints except 24

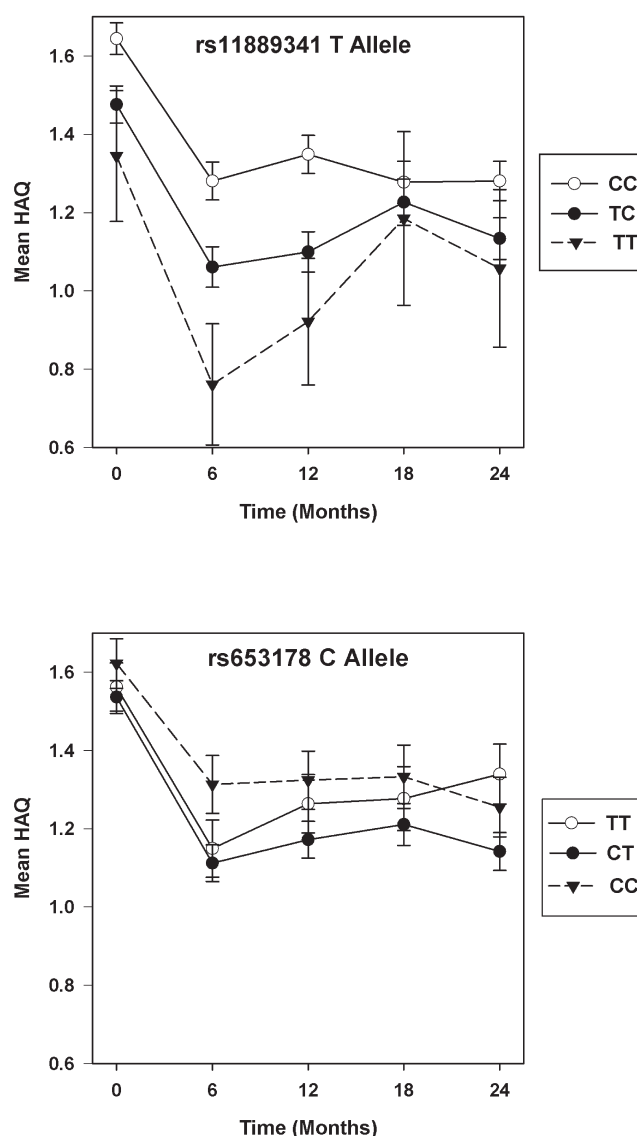


Figure 4. Significant genetic associations with HAQ scores over 6–24 months. Mean HAQ scores with standard error bars shown at each timepoint stratified by genotype. HAQ: Health Assessment Questionnaire.

months, mean HAQ scores were higher in CC homozygotes (Figure 4).

A significant association was observed between the full wGRS ( $\beta = -0.08$ ,  $p = 0.015$ ) and HLA wGRS ( $\beta = -0.05$ ,  $p = 0.026$ ) and latent growth factor means of HAQ score slope 1. The non-HLA wGRS did not associate with HAQ scores (Supplementary Table 7, available online at [jrheum.org](http://jrheum.org)).

## DISCUSSION

Our study represents a comprehensive analysis of the association between RA susceptibility variants and disease severity in early active RA, using LME and latent growth curve

modeling to evaluate whether risk variants predict radiograph, DAS28, and HAQ score progression. There are 3 main findings. The first and most important is that HLA susceptibility variants have a major influence on radiological progression with the main risk variants (\*04:01 and amino acid polymorphisms Val11 and His13 in HLA-DRβ1) and an HLA wGRS associating with greater joint destruction. By contrast, no association was observed between non-HLA susceptibility variants and radiographic progression when evaluating these variants individually and cumulatively using a wGRS. Our results suggest that non-HLA susceptibility loci do not have a clinically relevant influence on joint damage in early, active RA, because CARDERA was powered to detect markers accounting for 3.4% of radiographic progression variance and identified known joint destruction predictors (RF, ACPA, SE<sup>30,31</sup>).

The associations between \*04:01 and Val11 and radiographic damage have been reported in other RA populations, suggesting that the HLA associations in CARDERA are true-positive findings. Two large studies of established RA have reported associations between \*04:01 and erosive damage or cross-sectional radiographic scores<sup>30,32</sup>. The largest report — a metaanalysis of 29 studies (3240 patients) — demonstrated a significant association between the SE and erosions; a subanalysis of European studies with 4-digit HLA-DRB1 data available demonstrated this association was only significant for the \*04:01 allele. The effect of HLA-DRβ1 amino acid polymorphisms on radiographic damage has been reported in a conference abstract. In that analysis of patients with early RA from the Norfolk Arthritis Register (NOAR) and Early RA Study (ERAS), Val11 had a significant, SE-independent association with radiographic damage<sup>33,34</sup>. Although we were unable to establish whether Val11, His13, or the SE were driving their association with Larsen scores, when our results are considered alongside the NOAR/ERAS findings, it is likely that Val11 represents an independent predictor of joint destruction. This link has biological plausibility; Val11 is located in a peptide-binding groove and could mediate articular damage through influencing arthritogenic peptide presentation.

It is known that certain HLA variants predispose to ACPA<sup>35</sup>, which itself associates with erosive damage<sup>36</sup>. We found that \*04:01, His13, and Val11 only associated with radiographic damage in ACPA-positive RA. This suggests that their effects are not mediated by ACPA; if they caused joint destruction by driving ACPA formation, they would not associate with radiographic progression in ACPA-positive patients because they all harbor these autoantibodies. It also indicates that these 3 variants only predict radiographic progression in ACPA-positive disease. This finding has importance for 2 reasons. First, it suggests that patients at a substantially increased risk of radiographic progression can be identified by combining ACPA and HLA status. In CARDERA, ACPA-positive patients carrying 2 Val11 copies

had a mean ΔLarsen of 9.05 units (95% CI 5.94–12.15); ACPA-negative patients carrying no Val11 copies had a mean ΔLarsen of 2.77 units (95% CI 1.72–3.83). The clinical implication is that the former group may benefit from more intensive treatment than the latter group. Second, it indicates that as with gene-environment susceptibility factors<sup>19,37,38</sup>, ACPA-positive and ACPA-negative RA may also differ in their genetic risks for radiographic damage, further supporting the concept that they represent distinct disease entities.

Our second finding is that 2 SNP — rs11889341 (*STAT4*) and rs653178 (*SH2B3-PTPN11*) — associated with HAQ scores over 6–24 months. The relevance of these associations is, however, uncertain. Although the rs11889341 RA risk (T) allele associated with greater HAQ score increases at all timepoints, mean HAQ scores were lower in T allele carriers. Similarly, although the rs653178 RA risk (C) allele associated with improving HAQ scores at all timepoints except 24 months, mean HAQ scores were higher in CC homozygotes. Because of the uncertain clinical relevance of these associations, we did not seek to replicate them in alternative cohorts. A significant association was also seen between an HLA-wGRS and HAQ score over 0–6 months, although no associations were seen with individual HLA variants.

Our third finding is that RA susceptibility variants did not associate with disease activity. Although a significant association between an HLA-wGRS and baseline DAS28 was seen, this did not persist over time and no significant associations were seen with individual HLA-DRB1 alleles. To our knowledge, this is the first analysis of genetic associations with DAS28 after adjusting for treatment effects. Although several GWAS have attempted to identify genetic predictors of DAS28-defined anti-tumor necrosis factor responses<sup>18,39,40</sup>, only 1 definite association was identified (rs6427528 with etanercept response)<sup>18</sup>. One explanation for these largely negative findings is that DAS28 is not genetically determined. Heritability studies are needed to establish whether genetic markers could provide DAS28-defined prognostic data.

Our study has several unique features that strengthen existing evidence on genetic associations with RA severity. First, we examined genetic associations across many RA outcomes; previous studies have focused on radiographic damage alone<sup>2,3,5,6,7,41</sup>. Second, we evaluated genetic predictors in patients with early, active RA. Because the evidence supporting intensive combination therapy comes from RCT in such patients, our positive HLA findings have potential relevance to guiding decisions on treatment intensity. Third, to our knowledge, it is the first study of genetic associations with RA outcomes in an RCT; the standardized patient assessments increase the reliability of our findings. Finally, it evaluates all European RA susceptibility variants identified to date.

Our study has several other strengths. First, it was well



powered to detect clinically relevant predictors of each outcome. Second, the inclusion of individuals with severely active disease meant any findings are relevant to patient populations in whom current UK RA guidelines are focused<sup>42</sup>. Third, the randomization to treatment groups ensured treatment effects could be adjusted for. Fourth, high levels of radiographic progression were observed (39% of patients had clinically relevant 2-yearly Larsen score increases of  $\geq 4$  units that is comparable to observational study data in early RA<sup>43</sup>), further optimizing the study's power. Finally, our use of repeated-measures latent GCM meant we evaluated predictors of DAS28/HAQ scores at specific timepoints in an individual's disease course; in contrast, previous studies of DAS28/HAQ score predictors have mainly evaluated associations with these outcomes only at the study endpoint<sup>44,45,46,47</sup>.

There are also several limitations to our study. First, we were unable to exclude genetic associations of a small effect size, with our sample size being substantially smaller than the GWAS metaanalyses undertaken to identify susceptibility loci. However, such associations would be of too small an effect to influence clinical decision making. Second, disease outcomes were only evaluated over 24 months; genetic variants could exert their effects over longer time periods. Third, we did not replicate previously validated associations between *TRAF1*<sup>48</sup>, *IL2RA*<sup>6</sup>, and *C5orf30*<sup>49</sup> and radiographic progression. One probable explanation for this is that CARDERA is composed of patients with early, active RA, in contrast to the observational studies reporting these associations. Fourth, there was evidence of a modest correlation between DAS28 and HAQ. The Pearson *r* correlation coefficient comparing the 24-month change in DAS28 and HAQ scores was 0.51 ( $p < 0.001$ ); to a certain extent, the analyses of genetic associations with DAS28 and HAQ were therefore not independent. Finally, imputed HLA data were used that could not be internally validated (because of an absence of tissue typing data). However, the imputation method used (SNP2HLA) has been extensively validated<sup>16,50</sup> and has documented high accuracy at imputing ImmunoChip genotypes (97% and 99% at imputing 4-digit HLA alleles and amino acids, respectively) in 918 individuals from the 1958 Birth Cohort<sup>16</sup>. Additionally, \*04:01, His13, and Val11 had high INFO scores (0.8764, 0.9925, and 0.9810, respectively), indicating that their imputation was robust.

Further work is required to characterize the genetic architecture of radiological progression in RA. Our study has demonstrated that while HLA susceptibility variants have an important effect on joint destruction, non-HLA susceptibility variants do not associate with this trait in early, active RA. We consider that further candidate gene studies are unlikely to adequately characterize this trait and the metaanalysis of GWAS of radiographic progression is required, as undertaken for RA susceptibility. With large sample sizes, such an analysis would have sufficient power to detect salient

non-HLA variants and to determine which HLA amino acids have SE-independent effects. Although any identified loci are likely to have a small effect size, if used in combination with known clinical predictors such as ACPA, they could deliver clinically meaningful prognostic data, enabling stratified medicine in RA. Additionally, defining which loci drive articular damage could identify novel therapeutic targets facilitating drug discovery.

## ONLINE SUPPLEMENT

Supplementary data for this article are available online at [jrheum.org](http://jrheum.org).

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